

# A Yaglom type theorem for a branching process in Markovian environment

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# Gulf of Morbihan



# Outline

- 1 Formulation of the model.
- 2 Main results.
- 3 Intermediate results.
- 4 Elements of the proofs.

# What is the talk about

We deal with a **branching process** whose behaviour depend on a sequence called **environnement**.

The environment can be defined by

- an i.i.d. sequence of r.v.'s with values in some abstract space
- a stationary sequence of r.v.'s
- a deterministic sequence (so called varying environment)

The goal:

- 1 We shall consider the case when the environment is defined by a Markov chain with values in a finite spate space.
- 2 We shall prove a Yaglom type theorem.

# The Markovian environment

The environment is given on  $(\Omega, \mathcal{F}, \mathbb{P})$  by a homogeneous Markov chain  $(X_n)_{n \geq 0}$  with values in the finite state space  $\mathbb{X}$  and with the transition probabilities  $P(i, j)$ ,  $i, j \in \mathbb{X}$ .

- 1 The corresponding transition operator is denoted by  $\mathbf{P}$ .  $\mathbb{P}_i$  is the probability law on  $\mathbb{X}^{\mathbb{N}}$  and  $\mathbb{E}_i$  the associated expectation, for a starting point  $i \in \mathbb{X}$ .
- 2 We suppose the following: **The Markov chain  $(X_n)_{n \geq 0}$  is irreducible and aperiodic.**  $\implies$  the matrix  $\mathbf{P}^{k_0} > 0$  for some  $k_0$  (is primitive).
- 3 Then, by the Perron-Frobenius theorem we have the **spectral gap property**: there exist positive constants  $c_1, c_2$ , a unique positive  $\mathbf{P}$ -invariant probability  $\nu$  on  $\mathbb{X}$  ( $\nu(\mathbf{P}) = \nu$ ) and an operator  $\mathbf{Q}$  on  $\mathcal{C}(\mathbb{X})$  such that, for any  $g \in \mathcal{C}(\mathbb{X})$  and  $n \geq 1, i \in \mathbb{X}$ ,

$$\mathbf{P}g(i) = \nu(g) + \mathbf{Q}(g)(i) \quad \text{and} \quad \|\mathbf{Q}^n(g)\|_{\infty} \leq c_1 e^{-c_2 n} \|g\|_{\infty}, \quad (1)$$

where  $\mathbf{Q}(1) = 0$  and  $\nu(\mathbf{Q}(g)) = 0$  ( $\nu \perp \mathbf{Q}$ ).

# The branching process

The branching process  $(Z_n)_{n \geq 0}$  in the Markovian environment  $(X_n)_{n \geq 0}$  is defined as follows:

- 1 The initial population size is  $Z_0 = z \in \mathbb{N}$ .
- 2 For  $n \geq 1$ , we let  $Z_{n-1}$  be the population size at time  $n - 1$  and assume that at time  $n$ , provided the environment takes value  $X_n = i \in \mathbb{X}$ , the parent  $j \in \{1, \dots, Z_{n-1}\}$  generates  $\xi_i^{n,j}$  children.
- 3 Then the population size at time  $n$  is given by:

$$Z_n = \sum_{j=1}^{Z_{n-1}} \xi_{X_n}^{n,j}, \quad (2)$$

where the empty sum is equal to 0.

- 4  $(\xi_i^{n,j})_{j,n \geq 1}$  are i.i.d. with the same probability generating function

$$f_i(s) := \mathbb{E} \left( s^{\xi_i} \right), \quad s \in [0, 1]. \quad (3)$$

- 5  $\xi_i^{n,j}, j \geq 1, i \in \mathbb{X}$  are independent of  $X_0, \dots, X_n$  and  $Z_0, \dots, Z_n$

# Assumption

- 1 For any  $i \in \mathbb{X}$ , the random variable  $\xi_i$  satisfies the following:  
 $\mathbb{E}(\xi_i) > 0$  and  $\mathbb{E}(\xi_i^2) < +\infty$ .
  
- 2 Then, it follows that  $0 < f'_i(1) < +\infty$  and  $f''_i(1) < +\infty$ .

# The associated Markov walk

- 1 Introduce the function  $\rho : \mathbb{X} \mapsto \mathbb{R}$  satisfying

$$\rho(i) = \ln f'_i(1), \quad i \in \mathbb{X}.$$

- 2 Along with  $(Z_n)_{n \geq 0}$  consider the Markov walk  $(S_n)_{n \geq 0}$  such that  $S_0 = 0$  and, for  $n \geq 1$ ,

$$S_n = \ln (f'_{X_1}(1) \cdots f'_{X_n}(1)) = \sum_{k=1}^n \rho(X_k). \quad (4)$$



# Notation for the expectation

Consider the couple  $(X_n, Z_n)$ , where  $X_n$  represents the environment and  $Z_n$  the branching process.

- 1 The couple  $(X_n, Z_n)_{n \geq 0}$  is a Markov chain with values in  $\mathbb{X} \times \mathbb{N}$ ,
- 2 Let  $\mathbb{P}_{i,z}$  be the probability law on  $(\mathbb{X} \times \mathbb{N})^{\mathbb{N}}$  and  $\mathbb{E}_{i,z}$  the associated expectation generated by the finite dimensional distributions of the Markov chain  $(X_n, Z_n)_{n \geq 0}$  starting at  $X_0 = i$  and  $Z_0 = z$  ( $z$  initial particles).

## Non-lattice condition

- 1 The following non-lattice condition is needed to ensure that the local limit theorem for the Markov walk  $S_n$  holds true:

For any  $\theta, a \in \mathbb{R}$ , there exist  $n \geq 1$  and a path  $x_0, \dots, x_n$  in  $\mathbb{X}$  such that  $\mathbf{P}(x_0, x_1) \cdots \mathbf{P}(x_{n-1}, x_n) \mathbf{P}(x_n, x_0) > 0$  and

$$\rho(x_0) + \cdots + \rho(x_n) - (n+1)\theta \notin a\mathbb{Z}.$$

Recall that  $\rho(i) = \ln f'_i(1)$ ,  $i \in \mathbb{X}$ .

- 2 The above condition is equivalent to the condition that the Fourier transform operator

$$\mathbf{P}_{it}g(i) := \mathbf{P} \left( e^{it\rho} g \right) (i) = \mathbb{E}_i \left( e^{itS_1} g(X_1) \right), \quad g \in \mathcal{C}(\mathbb{X}), i \in \mathbb{X}, \quad (5)$$

has a spectral radius strictly less than 1 for  $t \neq 0$ .

Non-latticity for Markov chains with not necessarily finite state spaces is considered, for instance, in Shurenkov 1984 and Alsmeyer 1994.

# Spectral radius of the transfer operator

- 1 From the spectral gap property of the operator  $\mathbf{P}$  it follows that, for any  $\lambda \in \mathbb{R}$  and any  $i \in \mathbb{X}$ , the limit

$$k(\lambda) := \lim_{n \rightarrow +\infty} \mathbb{E}_i^{1/n} \left( e^{\lambda S_n} \right)$$

exists and does not depend on the initial state of the Markov chain  $X_0 = i$ .

- 2 Moreover, the number  $k(\lambda)$  is the spectral radius of the transfer operator  $\mathbf{P}_\lambda$ :

$$\mathbf{P}_\lambda g(i) := \mathbf{P} \left( e^{\lambda \rho} g \right) (i) = \mathbb{E}_i \left( e^{\lambda S_1} g(X_1) \right), \quad g \in \mathcal{C}(\mathbb{X}), i \in \mathbb{X}. \quad (6)$$

In particular, under the spectral gap and non-lattice conditions  $k(\lambda)$  is a simple eigenvalue of the operator  $\mathbf{P}_\lambda$  and there is no other eigenvalue of modulus  $k(\lambda)$ .

- 3 In addition, the function  $k(\lambda)$  is analytic on  $\mathbb{R}$ .

# Classification

- 1 The branching process in Markovian environment is said to be:
  - subcritical* if  $k'(0) < 0$ ;
  - critical* if  $k'(0) = 0$ ;
  - supercritical* if  $k'(0) > 0$ .

- 2 The following identity, has been established in Grama-Lauvergnat LePage 2018-SPA:

$$k'(0) = \nu(\rho) = \mathbb{E}_\nu(\rho(X_1)) = \mathbb{E}_\nu(\ln f'_{X_1}(1)) = \varphi'(0), \quad (7)$$

where  $\mathbb{E}_\nu$  is the expectation generated by the finite dimensional distributions of the Markov chain  $(X_n)_{n \geq 0}$  in the stationary regime and  $\varphi(\lambda) = \mathbb{E}_\nu(\exp\{\lambda \ln f'_{X_1}(1)\})$ ,  $\lambda \in \mathbb{R}$ .

- 3 Relation (7) proves that the classification made in the case of branching processes with Markovian environment and that for i.i.d. environment are coherent: when the random variables  $(X_n)_{n \geq 1}$  are i.i.d. with common law  $\nu$ , from (7) it follows that the two classifications coincide.

# Survival probability jointly with $X_n$

- 1 In the present paper we will focus on the critical case:

$$k'(0) = 0.$$

- 2 Our first result establishes the exact asymptotic of the survival probability of  $Z_n$  jointly with the event  $\{X_n = j\}$  when the branching process starts with  $z$  particles.

## Theorem 1

Assume that  $k'(0) = 0$ . Then, there exists a positive function  $u(i, z) : \mathbb{X} \times \mathbb{N} \mapsto \mathbb{R}_+^*$  such that for any  $(i, j) \in \mathbb{X}^2$  and  $z \in \mathbb{N}$ ,  $z \neq 0$ ,

$$\mathbb{P}_{i,z}(Z_n > 0, X_n = j) \underset{n \rightarrow +\infty}{\sim} \frac{u(i, z)\nu(j)}{\sqrt{n}}.$$

In the case  $z = 1$ , Theorem 1 has been proved in Grama Lauvergnat LePage 2018-SPA.

The proof for the case  $z > 1$  is not a direct consequence of the case  $z = 1$ .



## Additional condition

- 1 We shall complement the previous statement by studying the asymptotic behavior of  $Z_n$  given  $Z_n > 0$  under the following condition:

The random variables  $\xi_i, i \in \mathbb{X}$  satisfy:

$$\inf_{i \in \mathbb{X}} \mathbb{P}(\xi_i \geq 2) > 0.$$

- 2 The above condition is quite natural - it tells that each parent can generate more than 1 child with positive probability. In the present paper is used to prove the non-degeneracy of the limit of the martingale  $(\frac{Z_n}{e^{S_n}})_{n \geq 0}$ .

# Non-degeneracy of the limit law of the martingale $(\frac{Z_n}{e^{S_n}})_{n \geq 0}$

## Theorem 1

Assume  $k'(0) = 0$ . Then, for any  $i \in \mathbb{X}$ ,  $z \in \mathbb{N}$ ,  $z \neq 0$ , there exists a probability measure  $\mu_{i,z}$  on  $\mathbb{R}_+$  such that, for any continuity point  $t \geq 0$  of the distribution function  $\mu_{i,z}([0, \cdot])$  and  $j \in \mathbb{X}$ , it holds that

$$\lim_{n \rightarrow \infty} \sqrt{n} \mathbb{P}_{i,z} \left( \frac{Z_n}{e^{S_n}} \leq t, X_n = j, Z_n > 0 \right) = \mu_{i,z}([0, t]) \nu(j) u(i, z)$$

and

$$\lim_{n \rightarrow \infty} \mathbb{P}_{i,z} \left( \frac{Z_n}{e^{S_n}} \leq t, X_n = j \mid Z_n > 0 \right) = \mu_{i,z}([0, t]) \nu(j).$$

Moreover, it holds that  $\mu_{i,z}(\{0\}) = 0$ .

# The limit variance

- 1 Under adopted assumptions, the quantity

$$\sigma^2 := \nu(\rho^2) - \nu(\rho)^2 + 2 \sum_{n=1}^{+\infty} \left[ \nu(\rho \mathbf{P}^n \rho) - \nu(\rho)^2 \right] \quad (8)$$

is finite and positive, i.e.  $0 < \sigma < \infty$ .

See for example From Grama Lauvergnat Le Page 2017.



# Asymptotic of the normalized random walk $S_n$

1 Let

$$\Phi^+(t) = (1 - e^{-\frac{t^2}{2}})\mathbb{1}(t \geq 0), \quad t \in \mathbb{R},$$

be the Rayleigh distribution function.

2 The following assertion gives the asymptotic behavior of the normalized Markov walk  $S_n$  jointly with  $X_n$  provided  $Z_n > 0$ .

## Theorem 2

Assume  $k'(0) = 0$ . Then, for any  $i, j \in \mathbb{X}$ ,  $z \in \mathbb{N}$ ,  $z \neq 0$  and  $t \in \mathbb{R}$ ,

$$\lim_{n \rightarrow \infty} \sqrt{n} \mathbb{P}_{i,z} \left( \frac{S_n}{\sigma \sqrt{n}} \leq t, X_n = j, Z_n > 0 \right) = \Phi^+(t) \nu(j) u(i, z)$$

and

$$\lim_{n \rightarrow \infty} \mathbb{P}_{i,z} \left( \frac{S_n}{\sigma \sqrt{n}} \leq t, X_n = j | Z_n > 0 \right) = \Phi^+(t) \nu(j).$$

# The Yaglom type theorem

- 1 The following assertion is the Yaglom-type limit theorem for  $\log Z_n$  jointly with  $X_n$ .

## Theorem 3

Assume  $k'(0) = 0$ . Then, for any  $i \in \mathbb{X}$ ,  $z \in \mathbb{N}$ ,  $z \neq 0$ ,  $j \in \mathbb{X}$  and  $t \geq 0$ ,

$$\lim_{n \rightarrow \infty} \sqrt{n} \mathbb{P}_{i,z} \left( \frac{\log Z_n}{\sigma \sqrt{n}} \leq t, X_n = j, Z_n > 0 \right) = \Phi^+(t) \nu(j) u(i, z)$$

and

$$\lim_{n \rightarrow \infty} \mathbb{P}_{i,z} \left( \frac{\log Z_n}{\sigma \sqrt{n}} \leq t, X_n = j | Z_n > 0 \right) = \Phi^+(t) \nu(j).$$

## Intermediate results

For the proofs we need a number of intermediate results.

- 1 Introduce the first time when the Markov walk  $(y + S_n)_{n \geq 0}$  becomes non-positive: for any  $y \in \mathbb{R}$ , set

$$\tau_y := \inf \{k \geq 1 : y + S_k \leq 0\}. \quad (9)$$

### Theorem 4

Assume that  $k'(0) = 0$ .

1. For any  $(i, y) \in \mathbb{X} \times \mathbb{R}$  and  $j \in \mathbb{X}$ , we have

$$\lim_{n \rightarrow +\infty} \sqrt{n} \mathbb{P}_i(X_n = j, \tau_y > n) = \frac{2V(i, y)\nu(j)}{\sqrt{2\pi\sigma}}.$$

2. For any  $(i, y) \in \mathbb{X} \times \mathbb{R}$  and  $n \geq 1$ ,

$$\mathbb{P}_i(X_n = j, \tau_y > n) \leq c \frac{1 + \max(y, 0)}{\sqrt{n}}.$$

## Harmonic function $V$

- 1 The function  $V$  appearing in the previous theorem is the harmonic function (Doob transform) related to Markov walk  $(y + S_n)_{n \geq 0}$ . We state some of its properties to be used in the proofs.

### Proposition (Harmonic function)

There exists a non-negative function  $V$  on  $\mathbb{X} \times \mathbb{R}$  such that

1. For any  $(i, y) \in \mathbb{X} \times \mathbb{R}$  and  $n \geq 1$ ,

$$\mathbb{E}_i(V(X_n, y + S_n) ; \tau_y > n) = V(i, y).$$

2. For any  $i \in \mathbb{X}$ , the function  $V(i, \cdot)$  is non-decreasing and for any  $(i, y) \in \mathbb{X} \times \mathbb{R}$ ,

$$V(i, y) \leq c(1 + \max(y, 0)).$$

3. For any  $i \in \mathbb{X}$ ,  $y > 0$  and  $\delta \in (0, 1)$ ,

$$(1 - \delta)y - c_\delta \leq V(i, y) \leq (1 + \delta)y + c_\delta.$$

# Conditioned limit theorem

- 1 We need the asymptotic behaviour of the conditioned limit law of the Markov walk  $(y + S_n)_{n \geq 0}$  jointly with the Markov chain  $(X_n)_{n \geq 0}$ . (Extends previous results of the authors where the asymptotic of  $\frac{y+S_n}{\sigma\sqrt{n}}$  given the event  $\{\tau_y > n\}$  has been considered.)

## Theorem 5

Assume that  $k'(0) = 0$ . 1. For any  $(i, y) \in \text{supp}(V)$  and  $t \geq 0$ ,

$$\mathbb{P}_i \left( \frac{y + S_n}{\sigma\sqrt{n}} \leq t, X_n = j \mid \tau_y > n \right) \xrightarrow{n \rightarrow +\infty} \Phi^+(t)\nu(j).$$

2. There exists  $\varepsilon_0 > 0$  such that, for any  $\varepsilon \in (0, \varepsilon_0)$ ,  $n \geq 1$ ,  $t_0 > 0$ ,  $t \in [0, t_0]$  and  $(i, y) \in \mathbb{X} \times \mathbb{R}$ ,

$$\begin{aligned} \left| \mathbb{P}_i \left( \frac{y + S_n}{\sqrt{n}\sigma} \leq t, X_n = j, \tau_y > n \right) - \frac{2V(i, y)}{\sqrt{2\pi n}\sigma} \Phi^+(t)\nu(j) \right| \\ \leq c_{\varepsilon, t_0} \frac{(1 + \max(y, 0)^2)}{n^{1/2+\varepsilon}}. \end{aligned}$$

# Bound in the local limit theorem

- 1 We need a bound in the local limit theorem for the Markov walk  $(y + S_n)_{n \geq 0}$  jointly with the Markov chain  $(X_n)_{n \geq 0}$ .

## Theorem 6

Assume that  $k'(0) = 0$ . Then there exists  $c > 0$  such that for any  $a > 0$ , non-negative function  $\psi \in \mathcal{C}(\mathbb{X})$ ,  $y \in \mathbb{R}$ ,  $t \geq 0$  and  $n \geq 1$ ,

$$\begin{aligned} \sup_{i \in \mathbb{X}} \mathbb{E}_i(\psi(X_n) ; y + S_n \in [t, t + a], \tau_y > n) \\ \leq \frac{c \|\psi\|_\infty}{n^{3/2}} (1 + a^3) (1 + t) (1 + \max(y, 0)). \end{aligned}$$

# The dual Markov chain

- 1 Note that the invariant measure  $\nu$  is positive on  $\mathbb{X}$ . Therefore the dual Markov kernel

$$\mathbf{P}^*(i, j) = \frac{\nu(j)}{\nu(i)} \mathbf{P}(j, i), \quad i, j \in \mathbb{X} \quad (10)$$

is well defined.

- 2 On an extension of the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  we consider the dual Markov chain  $(X_n^*)_{n \geq 0}$  with values in  $\mathbb{X}$  and with transition probability  $\mathbf{P}^*$ . The dual chain  $(X_n^*)_{n \geq 0}$  can be chosen to be independent of the chain  $(X_n)_{n \geq 0}$ .

# The dual Markov walk

- 1 Accordingly, the dual Markov walk  $(S_n^*)_{n \geq 0}$  is defined by setting

$$S_0^* = 0 \quad \text{and} \quad S_n^* = - \sum_{k=1}^n \rho(X_k^*), \quad n \geq 1, \quad (11)$$

where we recall that  $\rho(i) = \ln f_i'(1)$ ,  $i \in \mathbb{X}$ .

- 2 For any  $y \in \mathbb{R}$  define the first time when the Markov walk  $(y + S_n^*)_{n \geq 0}$  becomes non-positive:

$$\tau_y^* := \inf \{k \geq 1 : y + S_k^* \leq 0\}. \quad (12)$$

For any  $i \in \mathbb{X}$ , denote by  $\mathbb{P}_i^*$  and  $\mathbb{E}_i^*$  the probability and the associated expectation generated by the finite dimensional distributions of the Markov chain  $(X_n^*)_{n \geq 0}$  starting at  $X_0^* = i$ .



# Duality

- 1 It can be verified that  $\nu$  is also  $\mathbf{P}^*$  invariant and that the transition operator  $\mathbf{P}^*$  satisfies the same condition as the transition operator  $\mathbf{P}$  of the forward Markov chain. This implies that the conditioned limit theorems hold also for the dual Markov chain  $(X_n^*)_{n \geq 0}$  and the Markov walk  $(y + S_n^*)_{n \geq 0}$ .
- 2 In particular there is a harmonic function  $V^*$  such that, for any  $(i, y) \in \mathbb{X} \times \mathbb{R}$  and  $n \geq 1$ ,

$$\mathbb{E}_i (V^*(X_n, y + S_n^*) ; \tau_y^* > n) = V^*(i, y).$$

- 3 The following duality property is obvious:

## Lemma 7

[Duality] For any  $n \geq 1$  and any function  $g: \mathbb{X}^n \rightarrow \mathbb{C}$ ,

$$\mathbb{E}_i (g(X_1, \dots, X_n) ; X_{n+1} = j) = \mathbb{E}_j^* (g(X_n^*, \dots, X_1^*) ; X_{n+1}^* = i) \frac{\nu(j)}{\nu(i)}.$$

# The change of probability measure

- 1 Fix any  $(i, y) \in \text{supp}(V)$  and  $z \in \mathbb{N}$ . The harmonic function  $V$  from Proposition (Harmonic function), allows us to introduce the probability measure  $\mathbb{P}_{i,y,z}^+$  on  $(\mathbb{X} \times \mathbb{N})^{\mathbb{N}}$  and the corresponding expectation  $\mathbb{E}_{i,y,z}^+$ , by the following relation: for any  $n \geq 1$  and any bounded measurable  $g: (\mathbb{X} \times \mathbb{N})^n \mapsto \mathbb{R}$ ,

$$\begin{aligned} & \mathbb{E}_{i,y,z}^+ (g(X_1, Z_1, \dots, X_n, Z_n)) \\ & := \frac{1}{V(i, y)} \mathbb{E}_{i,z} (g(X_1, Z_1, \dots, X_n, Z_n) \\ & \quad \times V(X_n, y + S_n) ; \tau_y > n). \end{aligned} \quad (13)$$

# Elements of the proof - 1

- 1 For any  $n \geq 1$ ,  $z \geq 1$  and  $s \in [0, 1]$ , the following quantity will play an important role in our study:

$$q_{n,z}(s) = 1 - (f_{X_1} \circ \dots \circ f_{X_n}(s))^z \in [0, 1].$$

For any  $n \geq 1$  and  $z \geq 1$ , the function  $s \mapsto q_{n,z}(s)$  is convex on  $[0, 1]$ . Since the sequence  $(\xi_i^{n,j})_{j,n \geq 1}$  is independent of the Markov chain  $(X_n)_{n \geq 0}$ , with  $s = 0$ , we have,  $\mathbb{P}_{i,y,z}^+$ -a.s.,

$$q_{n,z}(0) = \mathbb{P}_{i,y,z}^+(Z_n > 0 | (X_k)_{k \geq 0}). \quad (14)$$

## Elements of the proof - 2

- 1 The following formula is similar to the well-known statements from the papers by Agresti 1974 and Geiger and Kersting 2001: for any  $s \in [0, 1)$  and  $n \geq 1$ ,

$$\begin{aligned} \frac{1}{q_{n,z}(s)} &= \frac{1}{zf'_{X_1}(1) \cdots f'_{X_n}(1)(1-s)} \\ &+ \frac{1}{z} \sum_{k=1}^n \frac{\varphi_{X_k} \circ f_{X_{k+1}} \circ \cdots \circ f_{X_n}(s)}{f'_{X_1}(1) \cdots f'_{X_{k-1}}(1)} \\ &+ \psi_z \circ f_{X_1} \circ \cdots \circ f_{X_n}(s). \end{aligned} \tag{15}$$

## Elements of the proof - 3

- 1 We can rewrite (15) in the following more convenient form: for any  $s \in [0, 1)$  and  $n \geq 1$ ,

$$q_{n,z}(s)^{-1} = \frac{1}{z} \left( \frac{e^{-S_n}}{1-s} + \sum_{k=0}^{n-1} e^{-S_k} \eta_{k+1,n}(s) \right) + \psi_z \circ f_{X_1} \circ \dots \circ f_{X_n}(s), \quad (16)$$

where

$$\eta_{k,n}(s) = \varphi_{X_k} \circ f_{X_{k+1}} \circ \dots \circ f_{X_n}(s).$$

- 2 Since  $\frac{1}{2}\varphi(0) \leq \varphi(s) \leq 2\varphi(1)$ , for any  $k \in \{1, \dots, m\}$ ,

$$0 \leq \eta_{k,m}(s) \leq \frac{f''_{X_k}(1)}{f'_{X_k}(1)^2} \leq \eta := \max_{i \in \mathbb{X}} \frac{f''_i(1)}{f'_i(1)^2}. \quad (17)$$

# Non-degeneracy of martingale limit

We deal first with the non-degeneracy of the limit of the associated martingale  $\frac{Z_n}{e^{s_n}}$ . The next lemma establish the convergence of the Laplace transform.

## Lemma 8

Let  $s \geq 0$ . For any  $(i, 0) \in \text{supp} V$  and  $z \in \mathbb{N}$ ,  $z \neq 0$ , there exists a positive random variable  $W_{i,z}$  such that

$$\lim_{n \rightarrow \infty} \sqrt{n} \mathbb{E}_{i,z} (e^{-e^{-s} \frac{Z_n}{e^{s_n}}}; Z_n > 0, \tau_0 > n) = 2 \frac{V(i, 0)}{\sqrt{2\pi\sigma}} P_\infty(i, s, z),$$

where

$$P_\infty(i, s, z) := \mathbb{E}_{i,0,z}^+ (e^{-W_{i,z} e^{-s}}; \cap_{p \geq 1} \{Z_p > 0\}) \leq 1.$$

Moreover, for any  $(i, 0) \in \text{supp} V$  and  $z \in \mathbb{N}$ ,  $z \neq 0$ , it holds  $\mathbb{P}_{i,0,z}^+$ -a.s.

$$\cap_{p \geq 1} \{Z_p > 0\} = \{W_{i,z} > 0\}.$$

## Proof of lemma 8

- 1 For the proof we use a result from Kersting 2017 stated in a more general setting of branching processes with varying environment. To apply it we condition with respect to the environment  $(X_n)_{n \geq 0}$ , so that one can consider that the environment is fixed.
- 2 The condition (A) in Kersting 2017 is verified because we assumed that:

$$\inf_{i \in \mathbb{X}} \mathbb{P}(\xi_i \geq 2) > 0.$$

# The key point of the proof

The next result plays the key role in the proof.

## Theorem 9

For any  $i \in \mathbb{X}$ ,  $s \in \mathbb{R}$  and  $z \in \mathbb{N}$ ,  $z \neq 0$ ,

$$\begin{aligned} & \lim_{n \rightarrow \infty} \sqrt{n} \mathbb{E}_{i,z} \left( e^{-e^{-s} \frac{Z_n}{e^{S_n}}}; Z_n > 0 \right) \\ &= \frac{2}{\sqrt{2\pi\sigma}} \sum_{k=1}^{\infty} \mathbb{E}_{i,z} \left( V(X_k, 0) \mathbf{1}_{\text{supp} V(X_k, 0)} P_{\infty}(X_k, s + S_k, Z_k); \right. \end{aligned}$$

With  $s = +\infty$ ,  $Z_k > 0, T_k = k) =: u(i, z, e^{-s})$ .

where  $\lim_{n \rightarrow \infty} \sqrt{n} \mathbb{P}_{i,z}(Z_n > 0) = u(i, z)$ ,

$$u(i, z) = \frac{2}{\sqrt{2\pi\sigma}} \sum_{k=1}^{\infty} \mathbb{E}_{i,z} \left( V(X_k, 0) \mathbf{1}_{\text{supp} V(X_k, 0)} \mathbb{P}_{X_k, 0, Z_k}^+(W_{i,z} > 0); \right.$$

is from Theorem 1.  $Z_k > 0, T_k = k) > 0$



## Proof of Theorem 2

$$\lim_{n \rightarrow \infty} \sqrt{n} \mathbb{P}_{i,z} \left( \frac{S_n}{\sigma \sqrt{n}} \leq t, X_n = j, Z_n > 0 \right) = \Phi^+(t) \nu(j) u(i, z)$$

- 1 Let  $i, j \in \mathbb{X}$ ,  $y \geq 0$ ,  $z \in \mathbb{N}$ ,  $z \neq 0$ ,  $t \in \mathbb{R}$ . Then, for any  $n \geq 1$  and  $y \geq 0$ ,

$$\begin{aligned} & \mathbb{P}_{i,z} \left( \frac{S_n}{\sqrt{n}\sigma} \leq t, X_n = j, Z_n > 0 \right) \\ &= \mathbb{P}_{i,z} \left( \frac{S_n}{\sqrt{n}\sigma} \leq t, X_n = j, Z_n > 0, \tau_y > n \right) \\ &+ \mathbb{P}_{i,z} \left( \frac{S_n}{\sqrt{n}\sigma} \leq t, X_n = j, Z_n > 0, \tau_y \leq n \right) \\ &= l_1(n, y) + l_2(n, y). \end{aligned} \tag{18}$$

- 2 Using the bound in the conditioned limit theorem,

$$\sqrt{n} l_2(n, y) \leq \sqrt{ny} \mathbb{P}_{i,z}(Z_n > 0, \tau_y \leq n) \leq cze^{-y}(1+y) \xrightarrow{y \rightarrow \infty} 0. \tag{19}$$

## Proof of Theorem 2

- 1 To bound the first term  $I_1(n, y)$  we use the following lemma:

### Lemma

Suppose that  $i, j \in \mathbb{X}$ ,  $z \in \mathbb{N}$ ,  $z \neq 0$  and  $t \in \mathbb{R}$ . Then, for any  $y \geq 0$  sufficiently large,

$$\begin{aligned} \lim_{m \rightarrow \infty} \lim_{n \rightarrow \infty} \sqrt{n} \mathbb{P}_{i,z} \left( \frac{S_n}{\sqrt{n\sigma}} \leq t, X_n = j, Z_m > 0, \tau_y > n \right) \\ = \frac{2\Phi^+(t)}{\sqrt{2\pi\sigma}} V(i, y) U(i, y, z) \nu(j). \end{aligned} \quad (20)$$







# Proof of Yaglom theorem (Theorem 3)

- 1 Using the fact that  $\mu_{i,z}$  is a probability measure of mass 0 in 0 we show that

$$\lim_{n \rightarrow +\infty} \sqrt{n} \mathbb{P}_{i,z} \left( \left| \frac{\log Z_n}{\sigma \sqrt{n}} - \frac{S_n}{\sqrt{n}\sigma} \right| \geq \varepsilon, X_n = j, Z_n > 0 \right) = 0.$$

As  $\varepsilon$  is arbitrary, using Theorem 2 we conclude the proof.

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# Thank you !!!



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